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Research Article

Does Off-Farm Work Improve Farm Income? Empirical Evidence from Tolon District in Northern Ghana

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Participation of farm households in off-farm work has gained prominence in recent times as an income diversification strategy. The effect of off-farm work on farm income is however unclear. This paper therefore sought to provide empirical evidence of the income effect of off-farm activity participation using a cross section of maize farmers in Tolon District of Ghana as a case study. In order to account for sample selection bias, the Heckman selection model was used to estimate the factors influencing participation in off-farm work and the determinants of farm income. Furthermore, the study employed propensity score matching to evaluate the impact of off-farm work on farm income. The results indicate that participation in off-farm work is influenced by sex, age, and years of formal education of the respondent, farm size, and number of dependents while farm income is influenced by age of the respondent, farm size, and access to credit. In addition, the result of the propensity score matching revealed that participants in off-farm work increased their farm income by at least GH¢ 1702 as a result of income diversification. The rural economy therefore provides off-farm and on-farm linkages that enhance farmers' income from agriculture. The creation of employment opportunities outside the farm will therefore complement on-farm work and enhance income from farming.

1. Introduction

Most developing countries including Ghana depend on agriculture as an importance source of livelihood. It is estimated that in Ghana, more than 60 percent of the population are engaged in agriculture as a source of livelihood [1]. Majority of the farmers are smallholders who cultivate less than 2 hectares of farm land [2] and account for about 80 percent of the food produced locally [3]. Yields of most crops are generally below achievable levels due to reliance on natural rainfall for production, low adoption of improved production technologies, and lack of access to services such as agricultural extension and farm credit. In response to liquidity constraints and declining farm incomes, many smallholders diversify production and have multiple sources of income apart from farming, which have implications for agricultural productivity and farm income.

According to the existing literature, there is increasing recognition of the role that off-farm work plays particularly in smallholder agriculture in developing countries [4]. For most agrarian communities, farming is considered as the main occupation. Off-farm work is thus any activity undertaken by the farmer or farm household outside farming as an additional source of income. This is opposed to nonfarm work which relates to all activities that are not related to farming (such as dressmaking and commerce). Hence, offfarm work includes farm-related activities carried out by farm households for income such as exchange of labour for cash on another farmer's farm. The major sources of offfarm income in Ghana include commerce, agroprocessing, charcoal production, seasonal migration, brewing of local gin, basketry, and collection and sale of firewood [4, 5]. According to Chang and Wen [6], participation in off-farm work by farm households is a persistent phenomenon globally, with a steady increase in the dependence of farm

families on income from off-farm work. Income from off-farm work is regarded as an important source of livelihood for farm households and a means to diversify household income source. A study in rural Ghana by Jolliffe [7] indicated that about 74% of Ghanaian farm households participated in off-farm work. Also, research showed that, on average, 65% of American farm households were engaged in off-farm work [8]. Chang and Wen [6] also reported that about 75% of Taiwanese farm households earned off-farm salaries.

The increasing importance of off-farm work to farm households' economic well-being has generated a lot of discussion among researchers regarding the role off-farm work plays in household welfare, especially in terms of food security, agricultural productivity, and household income. It is a commonly held view that participation in off-farm work is expected to reduce on-farm labour availability and its allocation and thus constrain agricultural productivity. On the contrary, it is also believed that off-farm work enables farm households to stabilize household income and reduce vulnerability and uncertainties associated with agricultural production. As indicated by the extant literature, participation in off-farm work has two effects on production: a negative lost-labour effect and a positive income or liquidityrelaxing effect [9]. The lost-labour effect occurs when the household loses farm labour to off-farm activities, while the income effect occurs when the household earns income from off-farm activities which it can invest into farming. The effect of off-farm work on farm income will however depend on which of the two effects is stronger.

Off-farm work as a risk management tool that reduces income variability of farm households has been reported by some authors [10, 11]. As demonstrated by Mishra and Goodwin [12], farm households may depend on off-farm work to stabilize household income because farm commodity prices are more variable than off-farm wages. It is expected from the theory of production that a risk-averse farmer will choose to allocate labour and other resources to activities that are less risky to the point where the expected marginal returns are equal for the different activities. The higher variability in farm commodity prices is therefore expected to drive participation in off-farm work.

An earlier study by Lanjouw [13] noted that the rise in off-farm activity by farm households is as a result of declining farm incomes and the need to safeguard against production risks. This finding is corroborated by a recent study by Akinrinde et al. [14] which indicated that declining farm income is the main reason for income diversification among Nigerian farmers. Alasia et al. [15] on the other hand view participation in off-farm work as a form of self-insurance which enables farm households to stabilize household income.

Several studies allude to a positive effect of off-farm work on agricultural productivity, food security, and household income of farm households [16–18]. For most farm households, farm income constitutes a significant part of the total household income. A policy question which is relevant to most rural farm households is whether or not off-farm work contributes directly to farm income. In other words,

does income from off-farm work ease the liquidity constraints of farm households enabling greater use of farm resources in production? To the extent that farmers are able to invest earnings from off-farm activity into their farm business, farm output and productivity are expected to grow and exert a positive influence on farm income.

This study is motivated by the lack of empirical evidence of the effect of off-farm work on farm income of smallholder farmers in Ghana. The objective of this paper is therefore to assess the contribution of off-farm work to farm income in Tolon District of Ghana. The paper contributes to the literature on income diversification and its effects on household welfare by empirically estimating the magnitude and direction of impact of off-farm work on farm income of smallholder farmers.

The rest of the paper is structured as follows. Section 2 describes the methods used in the study, which provides a background of the study area, sampling and data collection, and method of data analysis and data description. Section 3 presents the results and discussion of the major findings of the study. The conclusion and recommendations from the study are provided in Section 4.

2. Materials and Methods

2.1. Study Area and Sampling. The study was conducted in the Tolon District of the Northern Region of Ghana. The area is part of the northern savannah zone of Ghana and is well known for its agricultural production. The area experiences a single rainfall regime per annum and is known for the cultivation of crops such as rice, maize, and groundnut, which are important staples. The population of the district stood at 72,990 according to the 2010 Population and Housing Census (PHC). An estimated 92% of the population are engaged in agriculture.

Three communities, namely, Nyankpala, Dundo, and Gbushalagu, were selected for the study due to their agricultural potential. Fifty farmers were randomly selected from each community to give a total sample of 150 farmers who were interviewed face-to-face using a semistructured questionnaire. Cochran's formula for sample size determination indicated that the number of farmers selected for the study is a representative sample. In the absence of a well-defined sample frame, households were selected at random in each community by taking into account the distribution of households. Information solicited from farmers included individual, household and farm characteristics, production activities, production costs and returns, and access to production resources and services.

2.2. Empirical Models. Participation in off-farm work and the factors determining farm income were analysed empirically using the Heckman Selection Model, while a nonparametric treatment effect model was used to estimate the effect of off-farm work on farm income. The Heckman selection model can be estimated using either maximum likelihood or a two-step approach. This study adopted the maximum likelihood approach because it overcomes some

of the limitations associated with the two-step approach such as the possibility of the correlation factor to lie outside the prescribed range. Methodologically, endogenous switching regression is an alternative approach for estimating the income effect of off-farm work. However, the Heckman selection model was preferred because it provided a better fit of the data.

2.2.1. Heckman Selection Model. Estimation of the Heckman selection model entails the estimation of a selection equation (off-farm participation model) using either probit or logit model followed by estimation of an outcome equation (farm income model) using least squares regression while accounting for sample selection bias. The assignment of households into participants and nonparticipants in off-farm work is nonrandom. Hence, OLS regression will not provide consistent parameter estimates of the outcome equation. To address this problem, there is the need to construct a correction factor, otherwise known as the inverse Mill's ratio, which is appended to the outcome equation as an additional explanatory variable.

The choice to participate in off-farm work can be estimated using a probit model which is specified as follows:

$$Z_i^* = \gamma' x_i + u_i, \tag{1}$$

where Z_i^* is a latent variable which measures the probability that the i^{th} household participates in off-farm work, such that the observed variable $Z_i=1$ if the household participates in off-farm work and $Z_i=0$ if otherwise, x_i indicates a vector of exogenous factors influencing Z_i^* , and γ represents a vector of parameters to be estimated.

In the second stage analysis, the amount of farm income (Y_i) is regressed on a set of exogenous factors, w_i , for all situations where the selection equation equals one $(Z_i = 1)$, using the estimated inverse Mill's ratio as an additional explanatory variable.

Hence, given that $Z_i = 1$, we have the following:

$$Y_i = \beta' w_i + v_i, \tag{2}$$

where Y_i indicates the amount of farm income, w_i is a vector of variables influencing farm income, and β is a vector of parameters to be estimated.

The error terms u_i and v_i have bivariate normal distributions with zero means, standard deviation δ_u and δ_v , and correlation coefficient ρ . While Z_i and x_i are both observable for a random sample, Y_i is observed only when the household participates in off-farm work ($Z_i = 1$). The Heckman selection model is specified as follows [19]:

$$E(Y_{i} | Z_{i} = 1) = E(Y_{i} | Z_{i}^{*} > 0)$$

$$= E(Y_{i} | u_{i} > \gamma' x_{i})$$

$$= \beta' w_{i} + E(v_{i} | u_{i} > \gamma' x_{i})$$

$$= \beta' w_{i} + \beta \sigma_{\nu} \lambda_{i} (\alpha_{\nu}),$$
(3)

where $\lambda_i(\alpha_u)$ is the inverse Mill's ratio, which has the following specification:

$$\lambda_{i}(\alpha_{u}) = \frac{\varphi(\alpha_{u})}{1 - \varphi(\alpha_{u})} = \frac{\varphi(-\alpha_{u})}{\varphi(\alpha_{u})} = \frac{\varphi(\gamma' x_{i}/\alpha_{u})}{\varphi(\gamma' x_{i}/\alpha_{u})}, \tag{4}$$

where φ represents the normal density function while ϕ indicates the normal distribution function.

As indicated by Heckman [19], when there is sample selection bias, an ordinary least squares (OLS) regression without the inverse Mill's ratio will lead to inconsistent parameter estimates. Thus, for this study, OLS regression of Y_i on w_i without the correction factor or inverse Mill's ratio $(\lambda_i(\alpha_u))$ will lead to inconsistent β estimates. Consequently, the inverse Mill's ratio was included as an additional explanatory variable in the outcome equation (equation (2)).

The identification criterion requires that at least one variable which influences participation in off-farm work but not farm income is included in the model. The number of dependents was chosen as an exclusion variable. This is because the number of dependents has a direct influence on participation in off-farm work, particularly on the number of hours worked but does not directly influence farm income. The choice of number of dependents as an exclusion variable is supported by Anang [20] who used the dependency ratio as an exclusion variable in a study on the effect of off-farm work on agricultural productivity in northern Ghana.

2.2.2. Propensity Score Matching: Estimating the Effect of Off-Farm Work on Farm Income. Impact evaluation studies in the extant literature have relied on estimation of average treatment effects as direct measures of the impact of interventions in the agricultural and other sectors. The effect of an intervention or exposure on those who received the treatment is an important measure in impact evaluation studies. Thus, to quantify the effect of off-farm work on farm income, we estimated the average treatment effect on the treated (ATT). The average treatment effect (ATE) given the observable data is denoted by

ATE =
$$E(Y^1 \mid T = 1) - E(Y^0 \mid T = 0),$$
 (5)

where Y^1 is the farm income of individuals who participated in off-farm work and Y^0 is the farm income of nonparticipants in off-farm work, T=1 represents the farmers who participated in off-farm work (referred to as the treated), and T=0 represents nonparticipants in off-farm work (referred to as the untreated or control). According to [21], $E(Y^1 \mid T=1) - E(Y^0 \mid T=0)$ is equal to zero for the case of a randomised design (i.e., in the absence of selection bias). However, in the presence of selection bias, the ATE result from equation (5) provides a biased estimate of the impact of off-farm work on farm income. Hence to overcome this bias, we need to estimate the average treatment effect on the treated (ATT), using the observational data, and conditioning on a vector of farm and household characteristics X as follows:

ATT =
$$E(\Delta | X, T = 1) = E(Y^{1} - Y^{0} | X, T = 1) = E$$

 $(Y^{1} | X, T = 1) - E(Y^{0} | X, T = 1).$ (6)

However, the counterfactual $E(Y^0 \mid X, T = 1)$ is unobservable, hence assumptions are made to estimate it as follows: $E(Y^0 \mid X, T = 1) = E(Y^0 \mid X, T = 0)$. Thus, the ATT equation becomes

ATT =
$$E(\Delta | X, T = 1) = E(Y^{1} - Y^{0} | X, T = 1)$$

= $E(Y^{1} | X, T = 1) - E(Y^{0} | X, T = 0).$ (7)

2.3. Sampling Procedure and Sources of Data. The data for the study were collected from smallholder maize farmers in 3 communities in the Tolon District of Northern Ghana. Fifty farmers were selected from each community to give a total sample of 150. The communities and respondents were randomly selected. Face-to-face interviews were carried out with each respondent using semistructured questionnaire. The questionnaire contained both open- and close-ended questions. The questions covered farm, household and institutional factors, and off-farm and production decisions, among others.

2.4. Description and Expected Signs of the Explanatory Variables. The description and expected signs of the variables included in the analysis are provided in Table 1. Male farmers are expected to have higher farm income, but sex is hypothesized to have an indeterminate effect on participation in off-farm work. Also, older farmers are expected to be more experienced in farming and more endowed with production resources than younger farmers, which is expected to lead to a higher farm income. Older farmers, by virtue of family obligations and dependents, are expected to have higher participation in off-farm work. Education enhances employability and is therefore hypothesized to increase participation in off-farm work. At the same time, education enhances the human capital which is expected to improve farm income. Farmers with access to credit are expected to have lower participation in off-farm work since credit eases the financial burden of farm households. Credit also enhances farm production through acquisition of farm inputs and financing of farm operations, which is expected to enhance farm profits. Farmers with large farms are expected to be progressive farmers and better-off compared to those with small farm holdings, leading to less participation in off-farm work and higher farm income. In addition, a higher number of dependents implies greater financial burden on households hence higher likelihood to take part in off-farm work. However, the number of dependents is not expected to have any direct effect on farm income.

3. Results and Discussion

3.1. Characteristics of the Respondents. We first describe the characteristics of the sample as shown in Table 2. The respondents have an average age of 38 years and 6.5 years of formal education. The respondents are therefore in their youthful age, a situation which is conducive for agricultural production. This is because agriculture in most developing countries, including Ghana, involves a lot of drudgery. The

respondents however have low level of education, a situation which can negatively affect uptake of innovation and ability to access and use information for agricultural production. Also, the respondents have an average household size of 9 and 3 dependent members and possess 2.7 hectares of land for agricultural production out of which 2 hectares is allocated to maize production. This shows that maize production is an important economic activity among rural dwellers. In Ghana, farm households depend on maize for food and income. Close to 43 percent of the respondents participated in off-farm work while 46 percent used credit in farming. In addition, 58 percent of the respondents are male indicating lower female participation in maize cultivation. Maize is a food security crop and household heads, most of whom are male, are anticipated to engage in its cultivation for home consumption and cash sales. On average, farmers' gross income from maize cultivation was GH¢ 2599 (approximately US\$ 490).

Table 3 presents the distribution of farmers' gross income from maize cultivation. Majority (76%) of the farmers obtained up to GH¢ 2500 as gross income from farming. The figure is very low, considering that maize is the most dominant crop grown by nearly every household in northern Ghana due to its important as a staple crop and also as cash crop for large scale producers. Participants in off-farm work had higher farm income compared with nonparticipants.

3.2. Results of the Heckman Selection Model. The results of the Heckman selection model are presented in Table 4. The log-likelihood ratio (LR) test is significant at 1% indicating the presence of selection bias, which justifies the use of the Heckman model to correct for the bias. In other words, estimating separate equations for the selection and outcome models would result in biased estimates.

3.2.1. Factors Influencing Participation in Off-Farm Activities. The estimates of the factors influencing participation in off-farm work using the Heckman selection model are presented in Table 4. The study indicates that the likelihood of maize farmers to participate in off-farm work is affected by age, sex, education, and total agricultural land.

Participation in off-farm work is higher for women. Women play several roles in the household apart from farming. They are involved in petty trading, arts, and craft and are more disposed to take up extra income earning activity to support the household's income. The marginal effect of sex indicates that the probability of women engaging in off-farm work is 0.214 higher than men. Ahmed and Melesse [22] as well as Man and Sadiya [23] obtained similar results in their studies in eastern Ethiopia and Malaysia, respectively. As observed by Ahmed and Melesse [22], female-headed households were more likely to engage in off-farm work compared with male-headed households because female-headed families take part in off-farm work to offset their relative lower farm income compared with maleheaded families. The result however disagrees with the findings of Pramanik et al. [24] in their study on the rural

TABLE 1. Description and	dexpected signs of t	he explanatory variables	included in the analysis.
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Variable	Description	Expected sign		
	Description	Model 1	Model 2	
Sex	Dummy = 1 if farmer is male; 0 otherwise	±	+	
Age	Age of farmer in years	+	+	
Education	Years of formal education	+	+	
Credit access	Dummy = 1 for credit access; 0 otherwise	_	+	
Maize farm size	Maize farm size in hectares	_	+	
Dependents	Number of dependents	+	NA	

¹ US\$ = 5.4 Ghana Cedis (GH¢). Model 1 is the probit participation model. Model 2 is the farm income (outcome) equation.

TABLE 2: Descriptive statistics of the variables used for the analysis.

Variable	Full s	Full sample		Participants		Nonparticipants	
	Mean	S. D.	Mean	S. D.	Mean	S. D.	Mean diff. [†]
Farm income	2599	3490	3457	4313	1962	2570	1495***
Sex	0.58	0.50	0.52	0.50	0.63	0.49	-0.11
Age	38.3	10.5	37.4	9.43	39.0	11.3	-1.64
Education	6.47	6.91	8.48	7.73	4.98	5.84	3.50***
Maize farm size	2.02	2.12	2.63	2.80	1.56	1.26	1.07***
Credit access	0.46	0.50	0.47	0.50	0.45	0.50	0.02
Dependents	2.80	2.31	2.17	2.11	3.27	2.36	-1.10***

[†]The *t* value of the difference in means between participants and nonparticipants. *Statistical significance at 10% level; **statistical significance at 1% level.

TABLE 3: Distribution of gross farm income of the respondents.

Variable	Full sa	Full sample		pants	Nonparticipants	
	Freq.	(%)	Freq.	(%)	Freq.	(%)
Up to 2,500	114	76.0	43	67.2	71	82.5
2,501-5,000	23	15.3	13	20.3	10	11.6
5,001-7,500	2	1.3	1	1.56	1	1.2
7,501-10,000	3	2.0	1	1.56	2	2.3
10,001-12,500	2	1.3	2	3.12	0	0
12,501-15,000	2	1.3	1	1.56	1	1.2
15,001-17,500	1	0.7	0	0	1	1.2
17,501-20,000	2	1.3	2	3.13	0	0
Above 20,000	1	0.7	1	1.56	0	0
Total	150	100	64	100	86	100
Mean	2,599		4,204		2,511	
Minimum	140		480		140	
Maximum	21,120		21,120		15,600	

nonfarm economy in Bangladesh as well as Beyene [25] in a study to assess the determinants of off-farm participation decision of farm households in Ethiopia.

Younger farmers are more likely to engage in off-farm work as shown by the coefficient of the age variable. As farmers become older, they become less disposed to engage in off-farm work. Younger people are more likely to find jobs because they are more energetic. Due to the low poverty level in most rural communities, off-farm employment for rural people involves labour-intensive activities which are better suited to younger people. The quadratic term for respondent's age shows that participation in work off-farm increases at a decreasing rate with age of the farmer. The result is consistent with Pramanik et al. [24] in their study in Bangladesh as well as Bila et al. [26] in their study on the

contribution of off-farm activities to farm income in Borno State, Nigeria.

Consistent with a priori expectation, the results indicate that educated farmers are more likely to work off-farm. Education enhances the human capital and opens up opportunities for employment off the farm. An additional year of education increases the likelihood of participation in off-farm work by 0.023. The results are consistent with Matshe and Young [27] in their study on off-farm labour allocation decisions in Zimbabwe, as well as Abdulai and CroleRees [28] in their study on income diversification among households in Southern Mali. The result also agrees with Seng [29] in a study on the effect of nonfarm work on household food consumption in rural Cambodia, Pramanik et al. [24] in their study on the rural nonfarm economy in

TABLE 4: Results of the	Heckman sel	ection model	of off-farm	work and fa	rm income
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Independent variables	Probit	model	Outcome	Outcome equation		
	Coefficient	Std. error	Coefficient	Std. error		
Sex	-0.523**	0.241	437.2	552.0		
Age	0.290***	0.096	-814.0***	223.9		
Age squared	-0.004***	0.001	12.26***	2.992		
Education	0.054***	0.018	-18.41	37.86		
Farm size	0.590***	0.127	849.6***	150.3		
Access to credit	-0.116	0.236	1281**	541.3		
Number of dependents	-0.121**	0.050				
Constant	-5.610***	1.901	14202***	4283		
Inverse Mill's ratio (λ)			-81.68*	0.095		
Rho	-0.881	0.079				
Sigma	2342.7	303.5				
Lambda	-2064.7	423.4				
Selected observations		ϵ	54	_		
Unselected observations		8	36			
Wald chi ² (6)		1	87			
Prob > chi ²		0.	00			

LR test of independent equations (rho = 0): chi²(1) = 8.94, prob > chi² = 0.003. *Statistical significance at 10% level. **Statistical significance at 5% level. **Statistical significance at 1% level. (6) and (1) indicate the degrees of freedom for the chi-squared estimation.

Bangladesh, and Owusu et al. [5] in their study on nonfarm work and food security in northern Ghana. McCarthy and Sun [30] estimated separate models for men and women and observed that household education levels had positive effect on both female and male participation in off-farm employment in rural northern Ghana.

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Participation in off-farm work was found to be positively related to farm size and significant at 1% level. This indicates that farmers with larger agricultural lands are more likely to take up employment off the farm. The result is contrary to a priori expectation and hard to explain. A unit increase in total agricultural land increases the likelihood of working off the farm by 0.041. The result agrees with Pramanik et al. [24] who found that farmers with larger plots were more likely to participate in nonfarm activities in Bangladesh. McCarthy and Sun [30] also observed that the size of owned land had a positive influence on participation in off-farm employment in rural northern Ghana.

The number of dependents had a negative and significant effect on participation in off-farm work, which is contrary to expectation. Households with many dependents are expected to have a higher propensity to participate in off-farm work, but the result suggests otherwise. The result is at variance with the findings of Anang [20] which indicated that an increase in the number of dependents increased the propensity of smallholder rice farmers to engage in off-farm work in northern Ghana.

3.2.2. Results of the Outcome (Farm Income) Model. The estimates of the parameters of the farm income equation using the Heckman selection model are presented in the 4th and 5th columns of Table 4. The results indicate that farm income initially decreases with age of the farmer as shown by the negative coefficient of the age variable. In other words, the younger the farmer, the higher the income from maize cultivation. However, as farmers become older and gain

experience in farming, their production and income increase. This is shown by the positive sign of the quadratic term of the age variable. As farmers gain experience in farming, this is expected to translate into more efficient ways of production and informed decision-making to maximise farm profits and income.

The results further indicate that the farm size variable is significant at 1% level and positively associated with farm income, which is consistent with a priori expectation. This implies that an increase in farm size results in an increase in farm income. Larger farm operators are therefore able to increase their income from farming. In addition, the study showed that the effect of credit on farm income was positive and significant at 10% level. Hence, access to credit enhanced the farm income of maize farmers in the study area. The result is consistent with a priori expectation as credit eases farm liquidity constraints and acquisition of farm inputs while enabling timeliness in carrying out farm operations to maximise output and profits.

3.3. Income Effect of Off-Farm Work. In order to quantify the income effect of off-farm work, we proceeded to estimate the average treatment effect on the treated (ATT) using the nearest neighbour and kernel matching methods (Table 5). We used the two matching methods in order to compare and check for robustness of the estimates. The results indicated that participation in off-farm work increases the farm income of participants by GH¢ 1702 using the nearest neighbour method and GH¢ 1776 in the case of the kernelbased method. The result indicates a positive and significant effect of off-farm work on farm income. What the result implies is that participants in off-farm work are able to increase their farm income by at least GH¢ 1702 as a result of income diversification. The result is supported by Osarfo et al. [31] who showed that participation in nonfarm work had a positive impact on the income of rural farm

TABLE 5: Estimates of the average treatment effect of off-farm work on farm inco
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Estimation method	Treated	Control	ATT	Robust S.E.	t statistic
Nearest neighbour matching	64	33	1701.7**	737.1	2.309
Kernel matching	64	79	1776.3***	526.3	3.375

Note: the outcome variable is farm income. **Statistical significance at 5% level. ***Statistical significance at 1% level.

households in the Upper East and Upper West Regions of Ghana. The important role that off-farm income plays in household income has been elaborated by other authors such as Ogbanje et al. [32] who estimated the off-farm income share of household income in North Central Nigeria at 50.28%.

4. Conclusion

The study assessed the effect of participation in off-farm activities on farm income of maize farmers in Tolon District of Northern Ghana. Due to the problem of sample selection bias, the study employed Heckman selection model to investigate the factors influencing participation in off-farm work and the determinants of farm income, while propensity score matching was used to estimate the impact of off-farm work on farm income. The results indicated that participation in off-farm work is influenced by sex, age, and years of formal education of the respondent, farm size, and number of dependents, while farm income is influenced by age of the respondent, farm size, and access to credit. Assessment of the impact of off-farm work on farm income indicated that participants in off-farm work increased their farm income by at least GH¢ 1702 as a result of income diversification. The rural economy therefore provides offfarm and on-farm linkages that enhance farmers' income from agriculture. The result also suggests that the negative lost-labour effect of off-farm activity participation is less than the positive liquidity (income) effect, resulting in an income gain for the farm. The study therefore concludes that participation in off-farm work enables maize farmers to improve their farm incomes, thereby improving household welfare. The creation of employment opportunities outside the farm will therefore complement on-farm work and enhance income from farming.

Data Availability

The data supporting the findings of the study are available upon request from the corresponding author.

Conflicts of Interest

The authors declare no conflicts of interest.

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